

Bivariate Multi-Fractal (BMF) Model

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Review of Multifractal Model

Multifractal Model of Asset Returns (MMAR).

Mandelbrot et al. (1974, 1997).

Returns $x(t)$ are assumed to follow a compound process:

$$x(t) = B_H[\theta(t)]$$

$B[\cdot]$: Fractional Brownian motion

H : Hurst Exponent

$\theta(t)$: Cumulative distribution function of a multi-fractal measure

$B[\cdot]$ and $\theta(t)$ are independent.

Combinatorial MF:

Review of Multifractal Model

Iterative Version MMAR:

Markov-switching Multi-Fractal (Calvet & Fisher, 2001):

$$x_t = \sigma \cdot \sqrt{\prod_{i=1}^k m_t^{(i)}} \cdot \varepsilon_t$$

Volatility components are renewed by transition probability γ_i

$$Prob(\text{new } m_t^{(i)}) = 1 - (1 - \gamma_1)^{(b^{k-1})}$$

⇒ Conceiving volatility as a hierarchical, multiplicative process with heterogeneous components.

Review of Multifractal Model

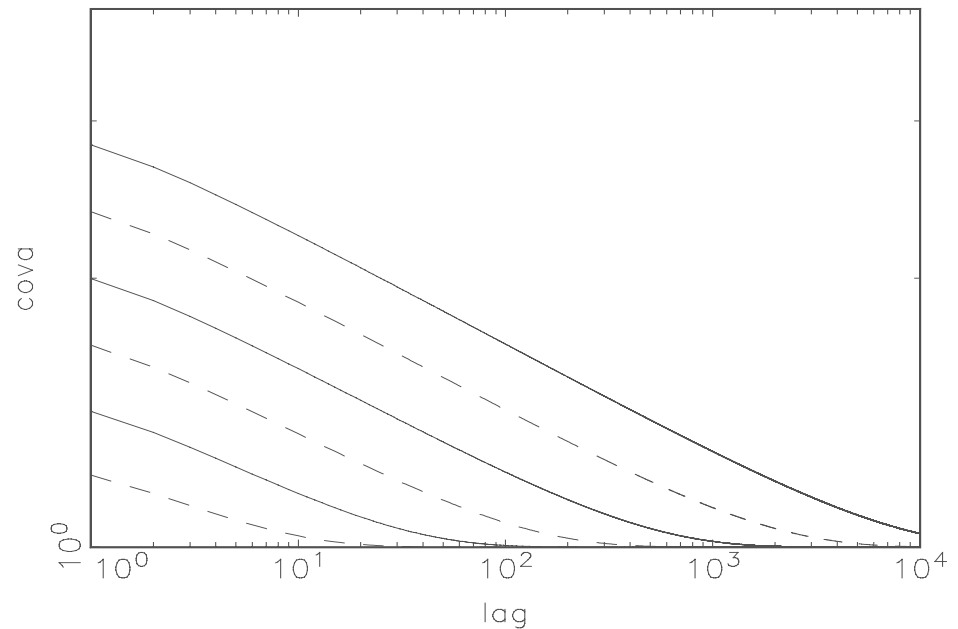
- Outliers.
- Volatility clustering.
- Asymptotic power-law behaviour of autocovariance function.

$$\text{Cov}(|x_t|^q, |x_{t+\tau}|^q) \propto \tau^{2d(q)-1}$$

Covariance of squared returns:

MF Binomial model:

$$\{1.3, 0.7\}, \quad k = 4, 6, \dots, 14$$



Bivariate Multifractal Model (BMF)

Two time series having certain amounts of joint cascade levels in both multifractal processes:

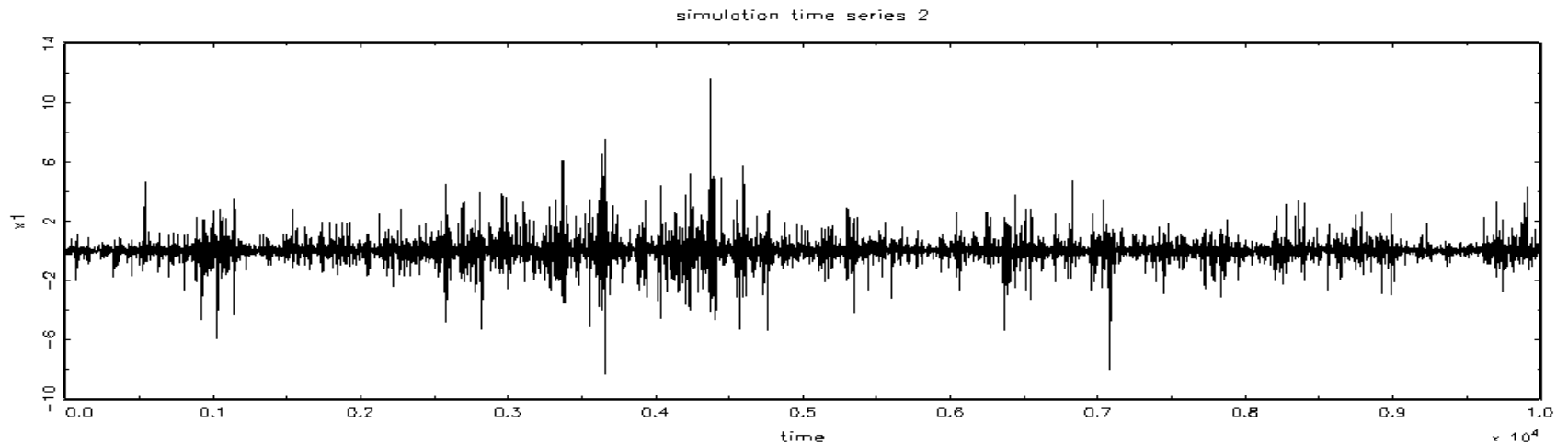
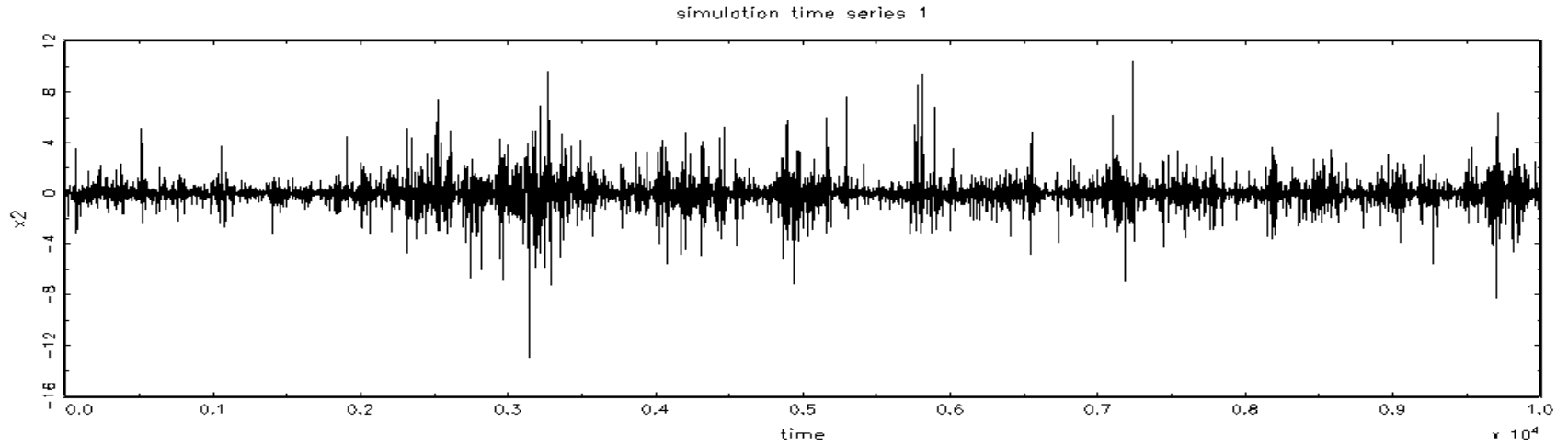
$$x_{t,q} = \sigma_q \cdot \left(\prod_{i=1}^k m_t^{(i)} \prod_{l=k+1}^n m_t^{(l)} \right)^{1/2} \cdot u_{t,q} \quad q = 1; 2, \quad u_{t,q} \sim BN(0, \Xi)$$

Transition probability:

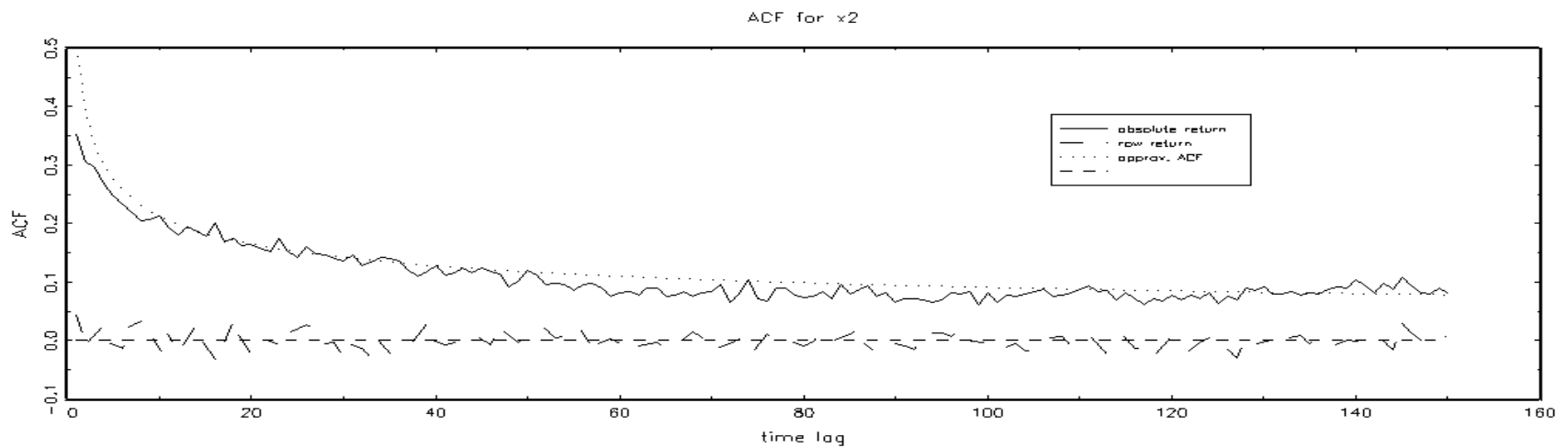
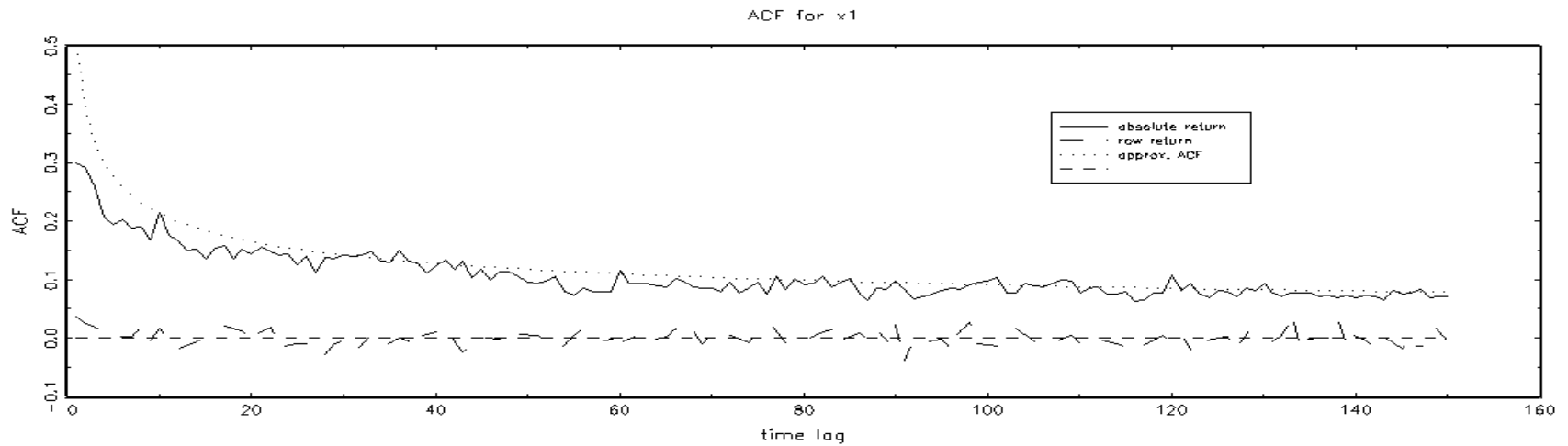
$$\gamma_j = 2^{-(k'-j)}$$

(depending on its rank within the hierarchical structure
or, remains unchanged with probability $1-r$)

Bivariate Multifractal Model



Bivariate Multifractal Model



Generalized Method of Moments

GMM (Hansen, 1982)

$$\hat{\beta} = \arg \min_{\beta \in \Theta} \bar{M}(\beta)' W \bar{M}(\beta)$$

Θ : parameter vector

$\bar{M}(\beta)$: vector of differences between sample moments
and analytical moments

W : positive definite weighting matrix

Generalized Method of Moments

Applicability of GMM on long memory models (Lux 2003, 2004)

Log increments transformation: $\varepsilon_t^{(i)} = \ln(m_t^{(i)})$

$$\begin{aligned} X_{t,\tau} &= \ln|x_t| - \ln|x_{t-\tau}| \\ &= 0.5 \sum_{i=1}^k \varepsilon_t^{(i)} + 0.5 \sum_{l=1}^n \varepsilon_t^{(l)} + \ln|u_t| - (0.5 \sum_{i=1}^k \varepsilon_{t-\tau}^{(i)} + 0.5 \sum_{l=1}^n \varepsilon_{t-\tau}^{(l)} + \ln|u_{t-\tau}|) \\ &= 0.5 \sum_{i=1}^k (\varepsilon_t^{(i)} - \varepsilon_{t-\tau}^{(i)}) + 0.5 \sum_{l=k+1}^n (\varepsilon_t^{(l)} - \varepsilon_{t-\tau}^{(l)}) + (\ln|u_t| - \ln|u_{t-\tau}|) \end{aligned}$$

Generalized Method of Moments

Moment conditions:

Group 1 --- Cross time series

$$Cov(X_{t,\tau}^q, Y_{t,\tau}^q)$$

$$Cov(X_{t,\tau}^q, Y_{t-\tau,\tau}^q)$$

Group 2 --- Individual time series

$$Cov(X_{t,\tau}^q, X_{t-\tau,\tau}^q)$$

$$Cov(Y_{t,\tau}^q, Y_{t-\tau,\tau}^q) \quad q = 1,2; \quad \tau = 1,5,10\dots$$

Generalized Method of Moments

Monte Carlo:

- Number of cascades: $n = 12$.
- Joint cascade levels: $k = 3, 6$.
- $\rho = 0.5$; $\sigma_1 = \sigma_2 = 1$.
 - (1) Binomial BMF: $\{m_0, 2-m_0\}$.
 - (2) Lognormal BMF: $-\log_2 M \sim N(\lambda, \sigma^2)$.
$$E[M] = 0.5, \rightarrow \sigma_M^2 = 2(\lambda - 1) / \ln 2.$$
- Population: $100,000$.
- Sample sizes: $N_1 = 2,000, N_2 = 5,000, N_3 = 10,000$.

GMM Binomial

		\hat{m}_0			$\hat{\rho}$			$\hat{\sigma}_1$			$\hat{\sigma}_2$		
		Bias	SE	RMSE	Bias	SE	RMSE	Bias	SE	RMSE	Bias	SE	RMSE
$m_0 = 1,20$	N1	-0.098	0.133	0.165	-0.002	0.071	0.071	-0.005	0.073	0.073	0.000	0.074	0.074
	N2	-0.081	0.120	0.145	0.001	0.045	0.045	0.000	0.046	0.046	0.000	0.046	0.046
	N3	-0.045	0.102	0.111	-0.001	0.030	0.030	0.000	0.033	0.033	0.003	0.034	0.034
$m_0 = 1,30$	N1	-0.111	0.149	0.186	0.009	0.072	0.072	-0.009	0.111	0.111	-0.002	0.112	0.112
	N2	-0.057	0.113	0.126	0.002	0.051	0.051	-0.001	0.070	0.070	0.000	0.069	0.069
	N3	-0.015	0.068	0.070	-0.003	0.033	0.033	0.000	0.051	0.051	0.004	0.051	0.051
$m_0 = 1,40$	N1	-0.083	0.140	0.162	0.012	0.078	0.079	-0.015	0.154	0.155	-0.006	0.153	0.153
	N2	-0.024	0.066	0.070	-0.002	0.052	0.052	-0.003	0.098	0.098	-0.001	0.094	0.094
	N3	-0.002	0.035	0.035	-0.008	0.033	0.034	0.000	0.071	0.071	0.004	0.069	0.069
$m_0 = 1,50$	N1	-0.042	0.088	0.097	0.002	0.086	0.086	-0.022	0.204	0.205	-0.014	0.197	0.198
	N2	-0.010	0.040	0.041	-0.010	0.055	0.056	-0.006	0.130	0.130	-0.004	0.123	0.123
	N3	0.003	0.024	0.025	-0.016	0.034	0.038	-0.001	0.095	0.095	0.003	0.091	0.091

GMM Binomial

		\hat{m}_0			$\hat{\rho}$			$\hat{\sigma}_1$			$\hat{\sigma}_2$		
		Bias	SE	RMSE	Bias	SE	RMSE	Bias	SE	RMSE	Bias	SE	RMSE
$m_0 = 1,20$	N1	-0.098	0.133	0.165	-0.002	0.071	0.071	-0.005	0.073	0.073	0.000	0.074	0.074
	N2	-0.081	0.120	0.145	0.001	0.045	0.045	0.000	0.046	0.046	0.000	0.046	0.046
	N3	-0.045	0.102	0.111	-0.001	0.030	0.030	0.000	0.033	0.033	0.003	0.034	0.034
$m_0 = 1,30$	N1	-0.111	0.149	0.186	0.009	0.072	0.072	-0.009	0.111	0.111	-0.002	0.112	0.112
	N2	-0.057	0.113	0.126	0.002	0.051	0.051	-0.001	0.070	0.070	0.000	0.069	0.069
	N3	-0.015	0.068	0.070	-0.003	0.033	0.033	0.000	0.051	0.051	0.004	0.051	0.051
$m_0 = 1,40$	N1	-0.083	0.140	0.162	0.012	0.078	0.079	-0.015	0.154	0.155	-0.006	0.153	0.153
	N2	-0.024	0.066	0.070	-0.002	0.052	0.052	-0.003	0.098	0.098	-0.001	0.094	0.094
	N3	-0.002	0.035	0.035	-0.008	0.033	0.034	0.000	0.071	0.071	0.004	0.069	0.069
$m_0 = 1,50$	N1	-0.042	0.088	0.097	0.002	0.086	0.086	-0.022	0.204	0.205	-0.014	0.197	0.198
	N2	-0.010	0.040	0.041	-0.010	0.055	0.056	-0.006	0.130	0.130	-0.004	0.123	0.123
	N3	0.003	0.024	0.025	-0.016	0.034	0.038	-0.001	0.095	0.095	0.003	0.091	0.091

GMM Lognormal

		\hat{m}_0			$\hat{\rho}$			$\hat{\sigma}_1$			$\hat{\sigma}_2$		
		Bias	SE	RMSE	Bias	SE	RMSE	Bias	SE	RMSE	Bias	SE	RMSE
$\lambda = 1,10$	N1	-0.029	0.052	0.060	0.011	0.078	0.079	-0.018	0.110	0.111	-0.004	0.115	0.115
	N2	-0.009	0.036	0.037	-0.001	0.049	0.049	-0.006	0.070	0.071	-0.001	0.072	0.072
	N3	-0.002	0.024	0.024	-0.005	0.039	0.039	-0.004	0.050	0.050	0.004	0.050	0.050
$\lambda = 1,20$	N1	-0.027	0.057	0.063	0.002	0.084	0.084	-0.029	0.164	0.166	-0.020	0.180	0.180
	N2	-0.005	0.037	0.038	-0.009	0.054	0.054	-0.017	0.115	0.116	-0.013	0.122	0.123
	N3	0.002	0.026	0.026	-0.015	0.036	0.039	-0.005	0.079	0.079	-0.009	0.085	0.086
$\lambda = 1,30$	N1	-0.022	0.063	0.067	-0.017	0.094	0.095	-0.042	0.246	0.250	-0.043	0.265	0.268
	N2	-0.005	0.036	0.037	-0.024	0.056	0.061	-0.026	0.183	0.185	-0.029	0.172	0.174
	N3	0.007	0.028	0.028	-0.031	0.039	0.050	-0.009	0.129	0.129	-0.016	0.121	0.122
$\lambda = 1,40$	N1	-0.024	0.068	0.072	-0.032	0.092	0.097	-0.047	0.339	0.341	-0.051	0.351	0.354
	N2	0.002	0.042	0.042	-0.042	0.067	0.079	-0.030	0.217	0.219	-0.038	0.221	0.224
	N3	0.010	0.029	0.031	-0.043	0.040	0.059	-0.023	0.203	0.204	-0.030	0.163	0.165

$$-\log_2 M \sim N(\lambda, \sigma^2), E[M] = 0.5, \rightarrow \sigma_M^2 = 2(\lambda - 1) / \ln 2$$

Maximum Likelihood Estimation

- The MF dynamics can be interpreted as a special case of Markov-switching process.
 - ⇒ Binomial case: Markov process with 4^k states for volatility
- The likelihood function can be derived by detecting the exact form of each possible component in the transition matrix.

Maximum Likelihood Estimation

Likelihood function :

$$\begin{aligned}
 L(x_1, \dots, x_T; \Theta) &= \prod_{t=1}^T f(x_t | x_1, \dots, x_{t-1}) \\
 &= \prod_{t=1}^T \left[\sum_{i=1}^{4^n} P(M_t = m^i | x_1, \dots, x_{t-1}) \cdot f(x) \right] \\
 &= \prod_{t=1}^T (\Omega_{t-1} \cdot A) f(x)
 \end{aligned}$$

$$\Omega_t^i = P(M_t = m^i | x_1, \dots, x_t)$$

A : Transition matrix $P(M_{t+1} = m^j | M_t = m^i)$

$f(x)$: Vector of conditional density $\{f(x_t | M_t = m^i)\}_i$

$$\Omega_t = \frac{f(x_t) \otimes (\Omega_{t-1} A)}{\sum f(x_t) \otimes (\Omega_{t-1} A)}$$

ML Binomial ($n = 5; k = 2$)

		Bias	SE	RMSE
$m_0 = 1.4$	N1	-0.001	0.023	0.023
	N2	0.000	0.020	0.019
	N3	0.003	0.019	0.019
$\rho = 0.5$	N1	0.053	0.094	0.106
	N2	0.043	0.092	0.100
	N3	0.040	0.084	0.091
$\sigma_1 = 1$	N1	0.053	0.093	0.106
	N2	0.042	0.091	0.099
	N3	0.040	0.083	0.091
$\sigma_2 = 1$	N1	-0.064	0.112	0.127
	N2	-0.053	0.107	0.118
	N3	-0.041	0.099	0.106

Simulated ML

Berzuini et al. (1996); Chib et al. (2002); Jacquier et al. (1994);
Pitt and Shephard (1999). ----- Particle filter:

$$\Omega_t^i \propto f(x_t | M_t = m^i) \sum_{j=1}^{4^n} P(M_t = m^i | M_{t-1} = m^j) \Omega_{t-1}^j$$

filtering

$$\approx f(x_t | M_t = m^i) \frac{1}{B} \sum_{b=1}^B P(M_t = m^i | M_{t-1} = m^{(b)})$$

Sampling/Importance Sampling (SIR):

$$P(q = b) = \frac{f(x_{t+1} | M_{t+1} = m_{t+1}^{(b)})}{\sum_{i=1}^B f(x_{t+1} | M_{t+1} = m_{t+1}^{(i)})}$$

particles

Computational convenient

Simulated ML

One-step ahead density:

$$\begin{aligned} f(x_t|x_{t-1}) &= \sum_{i=1}^{4^n} f(x_t|M_t = m^i)P(M_t = m^i|x_{t-1}) \\ &\approx \frac{1}{B} \sum_{b=1}^B P(x_t|M_t = m^{(b)}) \end{aligned}$$

Approximated Likelihood:

$$\hat{L}(x_1, \dots, x_T; \Theta) = \prod_{t=1}^T f(x_t|x_1, \dots, x_{t-1})$$

Value-at-Risk

Value-at-Risk (VaR):

- Specified target horizon
- Statistical confidence level
- The worst loss

Looking forward h -period return at time t : $\tilde{x}_{t,t+h} = \sum_{i=1}^h x_{t+i}$

VaR at the h -period horizon:

$$\Pr(\tilde{x}_{t,t+h} \leq VaR_{t,t+h}^{\alpha} | I_t) = \alpha$$

Value-at-Risk

Empirical daily data:

- Stock Exchange Index (*Jan. 1969 - Aug. 1999*)
 - Dow Jones Composite 65 Average Index*
 - Nikkei 225 Average Index*
- Foreign Exchange rates: (*Mar.1973 - Nov.2004*)
 - British Pound to US Dollar*
 - Australian Dollar to US Dollar.*
- Bonds: (*Jun.1976 - Oct.2004*).
 - U.S. 1 Year and 2 Year Treasury Constant Maturity Rate*

Returns: $x_t = \ln(p_t) - \ln(p_{t-1})$.

Failure Rate: Number of Observations above VaR.

Failure rate for multi-period VaR forecast (GMM)

		One day horizon			Two days horizon			Five days horizon		
		<i>DOW</i>	<i>NIK</i>	<i>EW</i>	<i>DOW</i>	<i>NIK</i>	<i>EW</i>	<i>DOW</i>	<i>NIK</i>	<i>EW</i>
<i>Stocks</i>	p = 10%	0.1074	0.0896	0.0974	0.1025	0.0912	0.1062	0.1052	0.0872	0.1110
	p = 5%	0.0625+	0.0422	0.0488	0.0627+	0.0451	0.0584	0.0604	0.0453	0.0645+
	p = 1%	0.0180	0.0081	0.0130	0.0193	0.0100	0.0165	0.0186	0.0087	0.0191
		<i>BP</i>	<i>AUD</i>	<i>EW</i>	<i>BP</i>	<i>AUD</i>	<i>EW</i>	<i>BP</i>	<i>AUD</i>	<i>EW</i>
<i>Fxs</i>	p = 10%	0.1059	0.9660	0.1104	0.1088	0.1035	0.1108	0.1051	0.0926	0.1051
	p = 5%	0.0542	0.0487	0.0512	0.0547	0.0501	0.0557	0.0563	0.0419	0.0469
	p = 1%	0.0106	0.0077	0.0085	0.0097	0.0092	0.0100	0.0118	0.0068	0.0069
		<i>T1</i>	<i>T2</i>	<i>EW</i>	<i>T1</i>	<i>T2</i>	<i>EW</i>	<i>T1</i>	<i>T2</i>	<i>EW</i>
<i>Bonds</i>	p = 10%	0.1010	0.1104	0.1083	0.1043	0.1096	0.1070	0.0902	0.1063	0.1041
	p = 5%	0.0577	0.0665+	0.0642+	0.0530	0.0659+	0.0662+	0.0541	0.0633+	0.0535
	p = 1%	0.0156	0.0187	0.0191	0.0146	0.0263+	0.0237+	0.0175	0.0229+	0.0227+

- EW denotes Equal-Weight portfolio
- + and * denote too risky and conservative VaR respectively

Failure rate for multi-period VaR forecast (ML)

		One day horizon			Two days horizon			Five days horizon		
		<i>DOW</i>	<i>NIK</i>	<i>EW</i>	<i>DOW</i>	<i>NIK</i>	<i>EW</i>	<i>DOW</i>	<i>NIK</i>	<i>EW</i>
<i>Stocks</i>	p = 10%	0.1175	0.0959	0.0904	0.1140	0.0968	0.0967	0.1125	0.0898	0.1000
	p = 5%	0.0690+	0.0439	0.0435	0.0677+	0.0465	0.0495	0.0650+	0.0475	0.0563
	p = 1%	0.0173	0.0063	0.0082	0.0180	0.0098	0.0120	0.0191	0.0088	0.0162
		<i>BP</i>	<i>AUD</i>	<i>EW</i>	<i>BP</i>	<i>AUD</i>	<i>EW</i>	<i>BP</i>	<i>AUD</i>	<i>EW</i>
<i>Fxs</i>	p = 10%	0.1176	0.1119	0.0903	0.1168	0.1143	0.0875	0.1095	0.0951	0.0892
	p = 5%	0.0577	0.0516	0.0403	0.0560	0.0517	0.0302*	0.0563	0.0431	0.0244*
	p = 1%	0.0084	0.0054	0.0031*	0.0082	0.0070	0.0027*	0.0106	0.0056	0.0031*
		<i>T1</i>	<i>T2</i>	<i>EW</i>	<i>T1</i>	<i>T2</i>	<i>EW</i>	<i>T1</i>	<i>T2</i>	<i>EW</i>
<i>Bonds</i>	p = 10%	0.1001	0.1092	0.1060	0.1020	0.1085	0.1041	0.0872	0.1019	0.0982
	p = 5%	0.0572	0.0662+	0.0609	0.0521	0.0648+	0.0653+	0.0527	0.0611	0.0528
	p = 1%	0.0156	0.0187	0.0191	0.0146	0.0263+	0.0237+	0.0175	0.0229+	0.0227+

- EW denotes Equal-Weight portfolio.
- + and * denote too risky and conservative VaR respectively

Conclusion (1)

MMF:

- Generalized Method of Moments Estimation.
 - No upbound for cascade levels.
 - Continuous distribution applicable.
- Maximum Likelihood Estimation.
 - Cascade levels bound: $n = 5$ for Bivariate Mf
and $n = 3$ for Trivariate MF.
 - Only for discrete distribution.

	Computation time (GAUSS)
GMM (Table 1)	2 Days
ML (Table 4)	5 Weeks

Conclusion (2)

Extension --- High Dimension MF:

$$x_{t,q} = \sigma_q \cdot \left(\prod_{i=1}^k m_t^{[i]} \prod_{h(\cdot)=k+1}^n m_t^{[h(\cdot)]} \right)^{1/2} \cdot u_{t,q}$$

σ_q and $u_{q,t}$ are q dimensional vectors

Estimation: GMM

Maximum Likelihood $(q^2)^n \times (q^2)^n$

SML